

# **Covariance and Sampling Distributions**

# Learning Outcomes

- Covariance
- Statistics and Inference
- Sampling Distributions
- Central Limit Theorem

# Covariance

# Covariance

The covariance measures the average dependence between multiple random variables. Let  $W = \begin{pmatrix} X \\ Y \end{pmatrix}$  be a random vector. The variance of  $W$  is defined as

$$\text{Var}(W) = \begin{pmatrix} \sigma_X^2 & \sigma_{XY} \\ \sigma_{XY} & \sigma_Y^2 \end{pmatrix}$$

Handwritten annotations in blue ink:

- $2 \times 2$  (written below the matrix)
- $2 \times 1$  (written above the matrix, with a line pointing to the first column)
- Covariances of RV's (written to the right of the matrix, with a line pointing to the off-diagonal elements)
- Variances of RV (written below the matrix, with a line pointing to the diagonal elements)

# Covariance

$\sigma_{xy}$

Let  $X_1$  and  $X_2$  be 2 random variables with mean  $\mu_1$  and  $\mu_2$ , respectively. The covariance of  $X_1$  and  $X_2$  is defined as

$$\begin{aligned} \text{Cov}(X_1, X_2) &= E\{(X_1 - \mu_1)(X_2 - \mu_2)\} \\ &= E(X_1 X_2) - \mu_1 \mu_2 \end{aligned}$$

If  $X_1$  and  $X_2$  are independent random variables, then

$$\text{Cov}(X_1, X_2) = 0$$

$$f(x, y) = \begin{cases} xy & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0 & \text{o.w.} \end{cases}$$

$$E(X) = \int_0^1 \int_0^1 x \cdot xy \, dx \, dy = \int_0^1 y \int_0^1 x^2 \, dx \, dy$$

$$E(Y) = \frac{1}{6}$$

$$\int_0^1 y \left[ \frac{x^3}{3} \right]_0^1 dy$$

$$\int_0^1 \frac{1}{3} y \, dy$$

$$E(XY) = \int_0^1 \int_0^1 xy \cdot xy \, dx \, dy$$

$$\frac{1}{3} \left[ \frac{y^2}{2} \right]_0^1 = \frac{1}{6}$$

$$\int_0^1 y^2 \int_0^1 x^2 \, dx \, dy$$

$$\int_0^1 \frac{1}{3} y^2 \, dy = \frac{1}{3} \left[ \frac{y^3}{3} \right]_0^1 = \frac{1}{9}$$

$$\text{Cov}(X, Y) = \frac{1}{9} - \frac{1}{6} \cdot \frac{1}{6} = \frac{1}{9} - \frac{1}{36} = \frac{4}{36} - \frac{1}{36} = \frac{3}{36}$$

# Correlation

The correlation of  $X_1$  and  $X_2$  is defined as

$$\rho = \text{Cor}(X_1, X_2) = \frac{\text{Cov}(X_1, X_2)}{\sqrt{\text{Var}(X_1)\text{Var}(X_2)}}$$



# MGF Property: Independence

Let  $X$  and  $Y$  be independent random variables. Let  $Z = X + Y$ , the MGF of  $Z$  is

$$M_Z(t) = M_X(t)M_Y(t)$$

$$Z = aX + bY$$

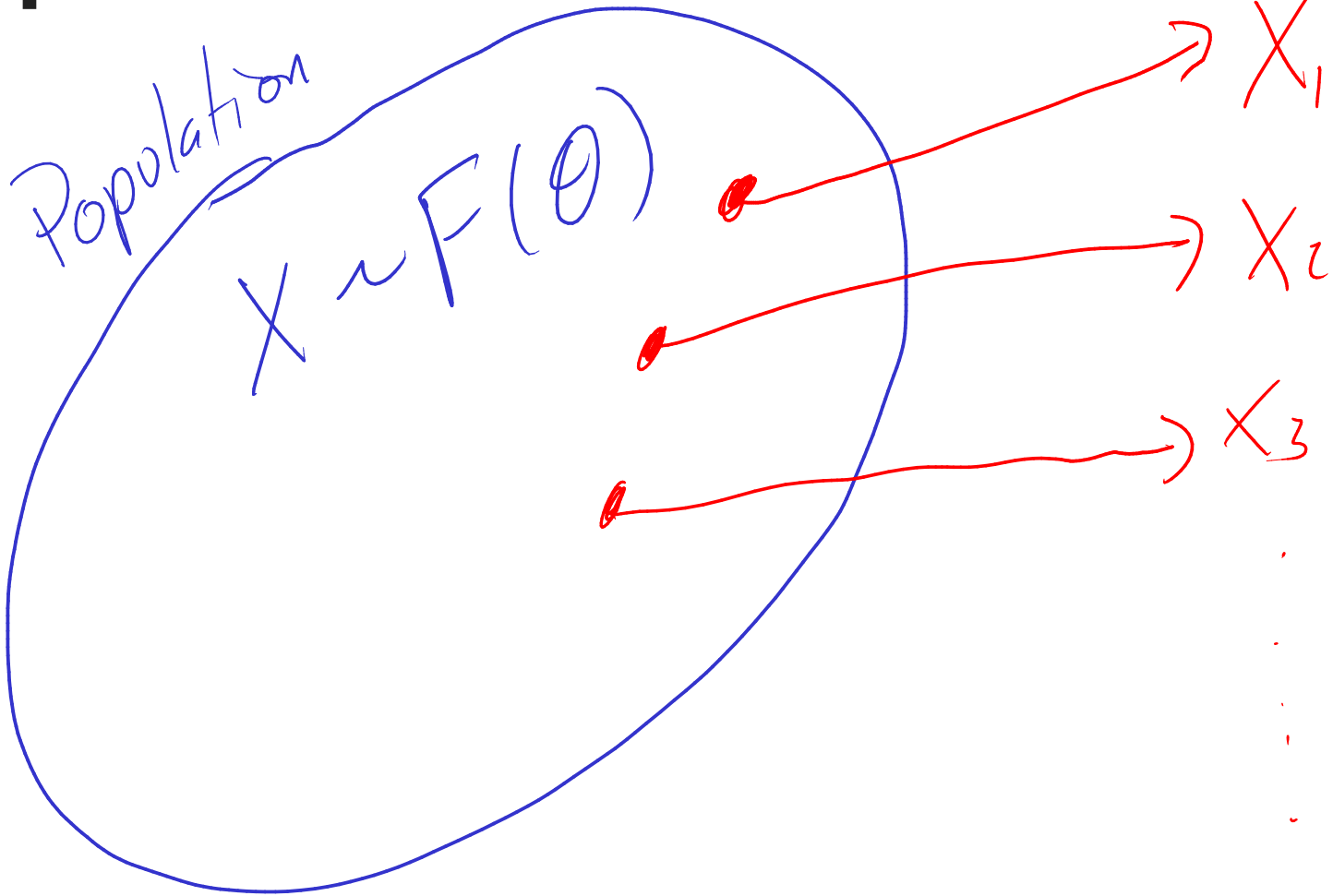
$$E(Z) = aE(X) + bE(Y)$$

$$\text{Var}(Z) = a^2 \text{Var}(X) + b^2 \text{Var}(Y)$$

# Statistics and Inference

# Sample

$$X_i \sim F(\theta)$$



$$\underline{X} = \{X_1, \dots, X_n\}$$

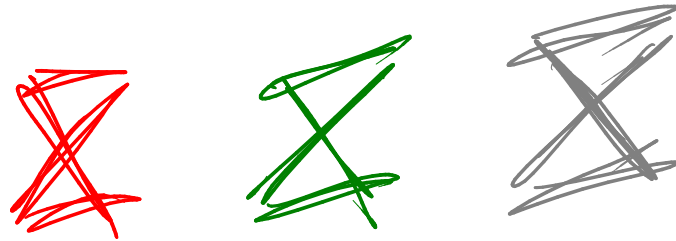
# Statistics

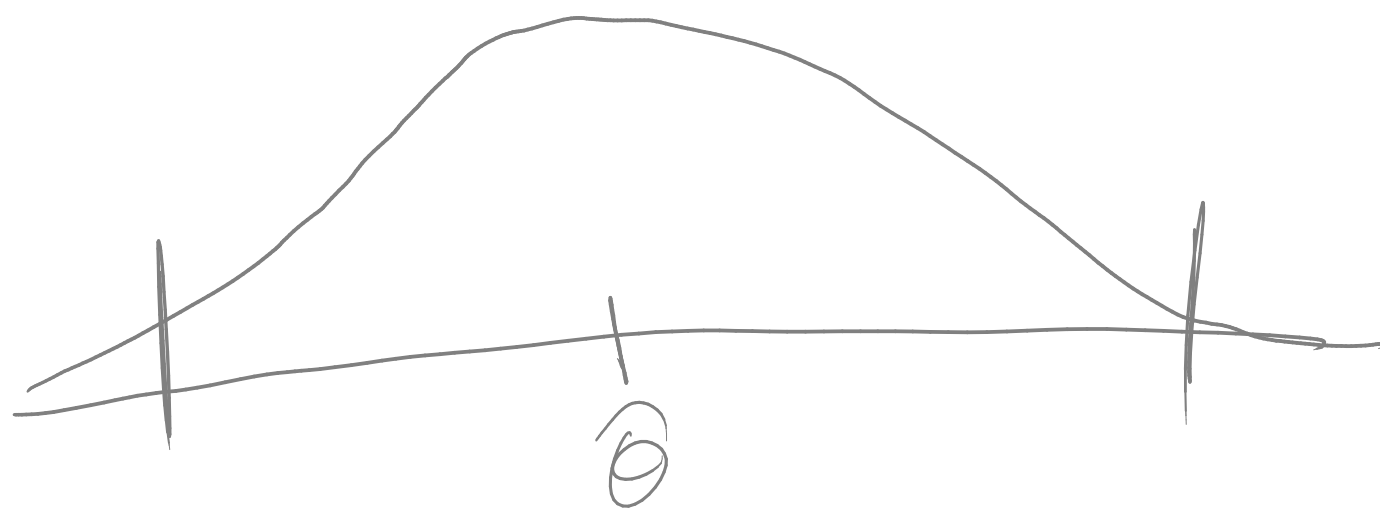
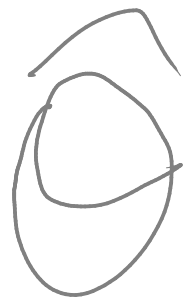


$$\underline{X} = \{X_1, \dots, X_n\}$$

hat estimate

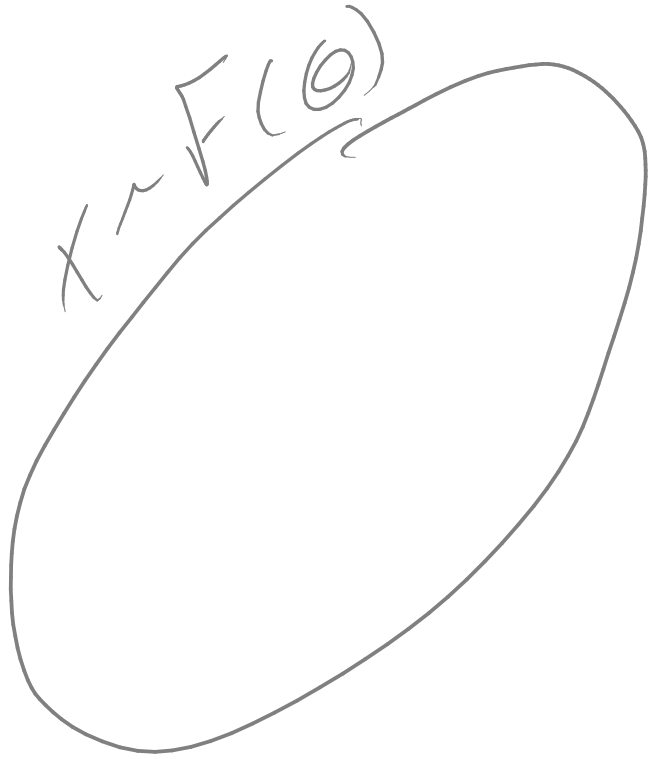
$$\hat{\theta} = g(\underline{X})$$





# Inference

$$\theta = 5$$



$$\text{Error} \\ \alpha = 0.011$$

$$\hat{\theta} = 4.8 \pm 0.1$$

$$\hat{\theta} = 4 \pm 1.2$$

$$\alpha = 0.001$$

# Sampling Distributions



# iid Random Variables

$$\underline{X} = \{X_1, \dots, X_n\}$$

$$X_i \perp X_j \quad \leftarrow \text{independent}$$

identical and independent

$$X_i \sim F(\theta)$$

# Sampling Distributions

A sampling distribution is the distribution of a statistic. Many known statistics have a known distribution.

$$\hat{\theta} \sim G(\theta)$$

$\bar{X}$ 

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

$$X_i \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2)$$

$$E(X_i) = \mu$$

$$\text{Var}(X_i) = \sigma^2$$

$$E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} E\left(\sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i)$$

$$= \frac{1}{n} \sum_{i=1}^n \mu = \frac{n\mu}{n} = \mu$$

$$\sum_{i=1}^5 E(X_i) = \underbrace{E(X_1)}_{\mu} + \underbrace{E(X_2)}_{\mu} + \underbrace{E(X_3)}_{\mu} + \underbrace{E(X_4)}_{\mu} + \underbrace{E(X_5)}_{\mu}$$

$$5\mu$$

$$\text{Var}(\bar{X}) = \text{Var}\left(\frac{1}{n} \sum X_i\right) = \frac{1}{n^2} \text{Var}\left(\sum X_i\right)$$

$$= \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) = \frac{1}{n^2} \sum_{i=1}^n \sigma^2 = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

$s^2$ 

Sample variance

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$$

$$s^2 \sim \frac{(n-1)}{\sigma^2} \chi^2(n-1)$$

# t-distribution

Let  $Z \sim N(0, 1)$ ,  $W \sim \chi^2_\nu$ ,  $Z \perp W$ ; therefore:

$$T = \frac{Z}{\sqrt{W/\nu}} \sim t_\nu$$

t-tests

# F-distribution

Let  $W_1 \sim \chi_{\nu_1}^2$ ,  $W_2 \sim \chi_{\nu_2}^2$ , and  $W_1 \perp W_2$ ; therefore:

$$F = \frac{W_1/\nu_1}{W_2/\nu_2} \sim F_{\nu_1, \nu_2}$$

# Example

Let  $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ , show that  $\bar{X} \sim N(\mu, \sigma^2/n)$ .

Note: the MGF of  $X_i$  is  $e^{\mu t + \frac{t^2 \sigma^2}{2}}$ .



# Central Limit Theorem

# Central Limit Theorem

Let  $X_1, X_2, \dots, X_n$  be identical and independent distributed random variables with  $E(X_i) = \mu$  and  $Var(X_i) = \sigma^2$ . We define

$$Y_n = \sqrt{n} \left( \frac{\bar{X} - \mu}{\sigma} \right) \text{ where } \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i.$$

Then, the distribution of the function  $Y_n$  converges to a standard normal distribution function as  $n \rightarrow \infty$ .

$$Y_n \sim N(0, 1)$$

# Central Limit Theorem

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

$$n \rightarrow \infty$$

# Example

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \chi_p^2$ , the MGF is  $M(t) = (1 - 2t)^{-p/2}$ . Find the distribution of  $\bar{X}$ .

$$E(x) = p \quad \text{Var}(x) = 2p = \sigma^2$$

$$n \rightarrow \infty$$

$$\bar{X} \sim \mathcal{N}\left(p, \frac{2p}{n}\right)$$



$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

Cauchy Dist

~~$F(x)$~~